

ECE 6540, Lecture 14

Bayesian Estimation:

LMMSE Estimators and Scalar Kalman Filters

Last Time

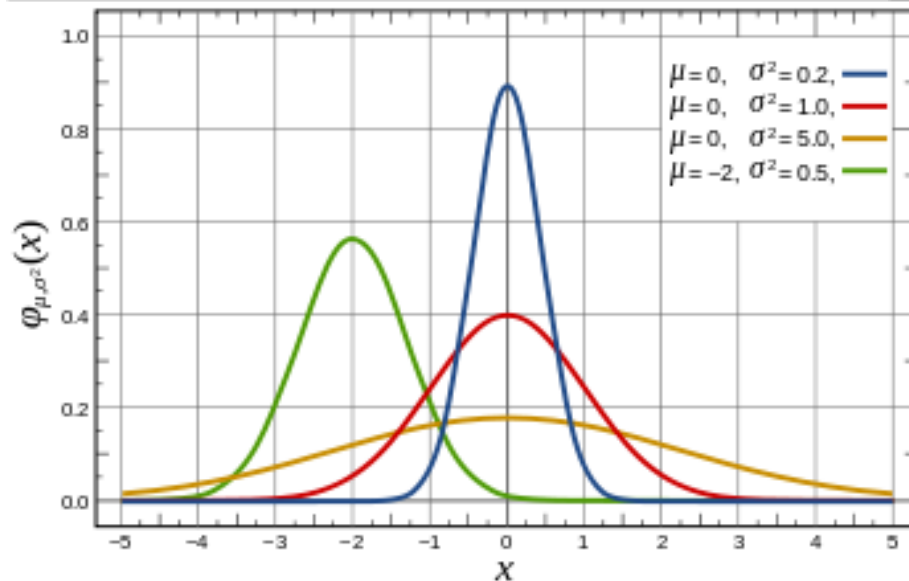
- Bayes Risk (Average Loss)
- Generalized Bayesian Estimator
- MMSE Estimator
- Maximum A Posterior (MAP) Estimator
- Linear MMSE Estimator

Example:
MAP Estimator

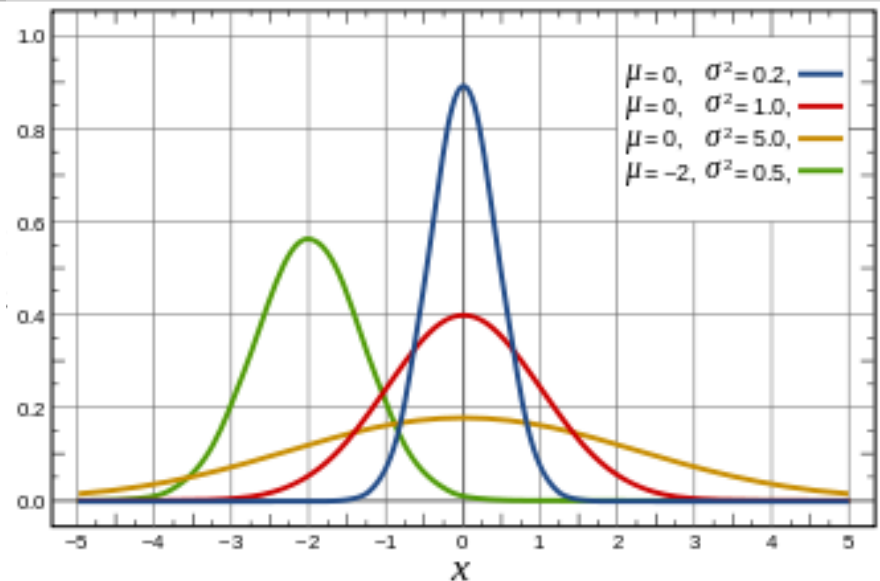
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■ MAP Estimator Example

- $x = H\theta + w$
- Gaussian likelihood $x | \theta \sim \mathcal{N}(\theta, \sigma^2 I)$
- Assume Gaussian prior $\theta \sim \mathcal{N}(0, \sigma_\theta^2 I)$
- Why might we choose this prior?



Likelihood



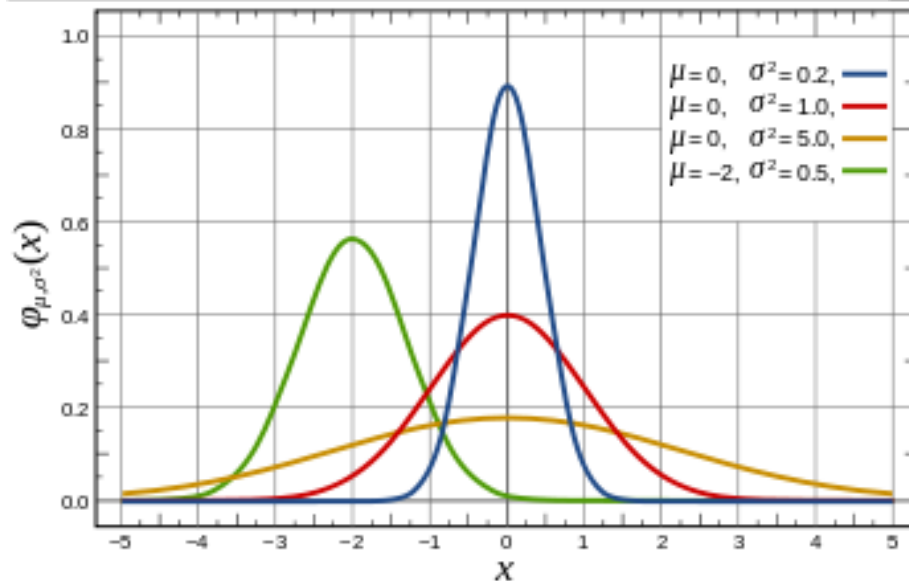
Prior

Last Time

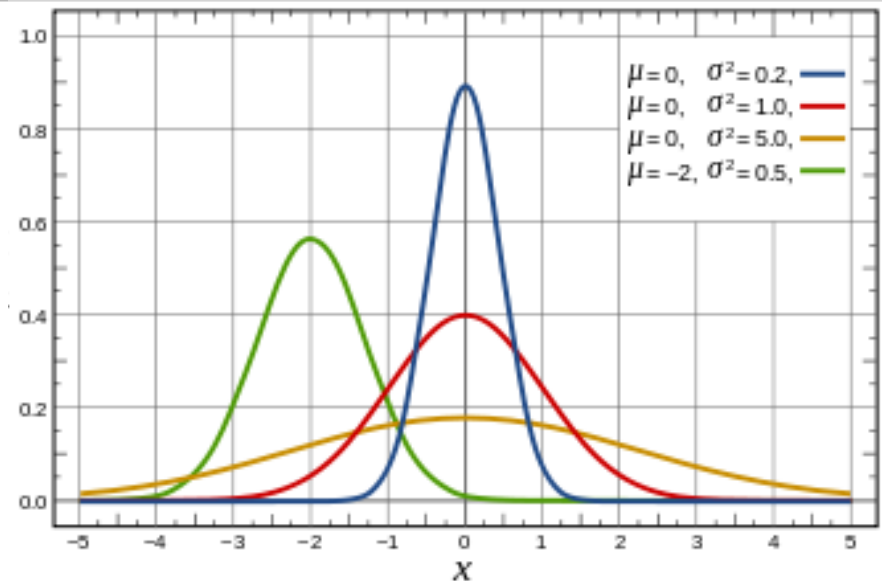
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Difference Operator



Likelihood

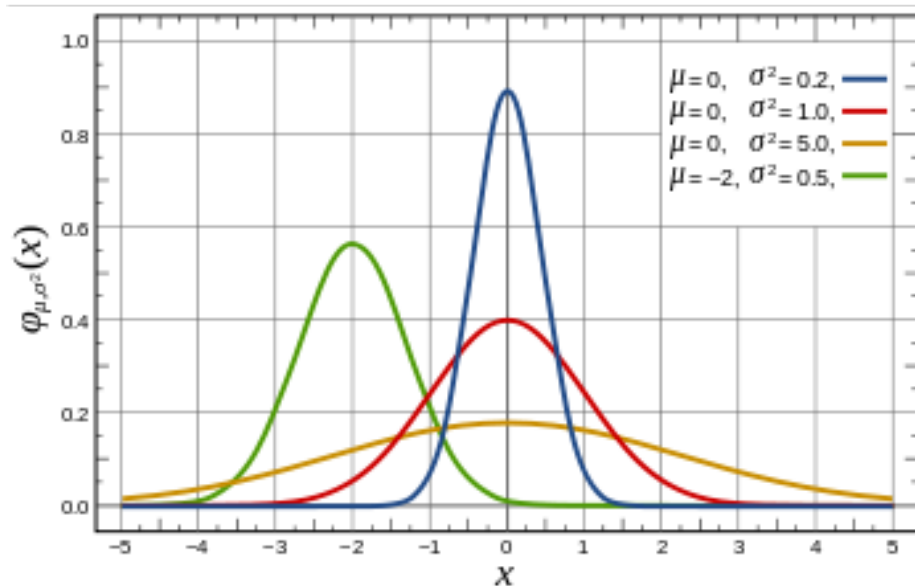


Prior

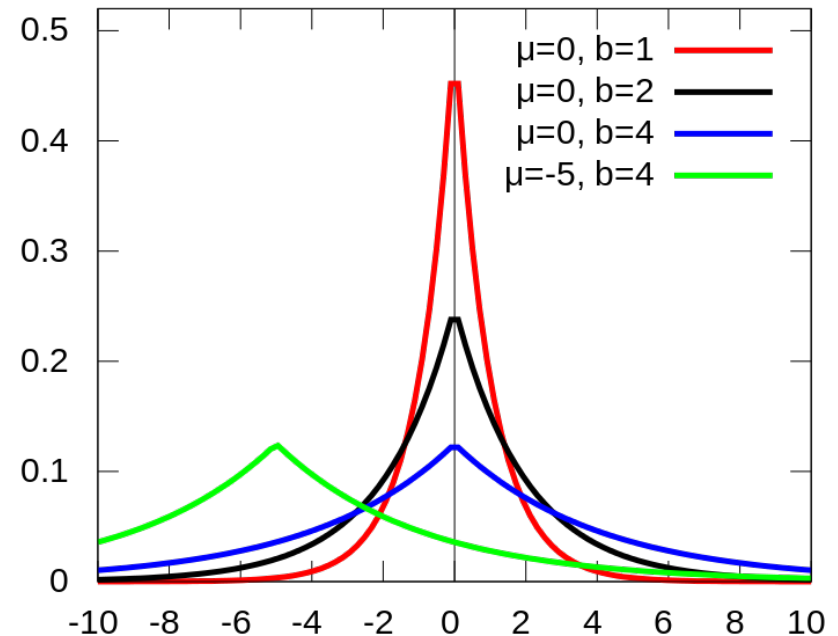
Last Time

■ MAP Estimator Example

- $x = H\theta + w$
- Gaussian likelihood $x | \theta \sim \mathcal{N}(\theta, \sigma^2 I)$
- Assume Laplace prior $\theta_n \sim \text{Laplace}(0, b)$
- **Why might we choose this prior?**



Likelihood



Prior

Source: https://en.wikipedia.org/wiki/Laplace_distribution#/media/File:Laplace_pdf_mod.svg

Source: https://en.wikipedia.org/wiki/Normal_distribution#/media/File:Normal_Distribution_PDF.svg

Last Time

■ MAP Estimator Example

- Gaussian likelihood $x | \theta \sim \mathcal{N}(0, \theta)$
- Assume inverse gamma prior $p(\theta) = \frac{\beta^\alpha}{\Gamma(\alpha)} \theta^{-\alpha-1} \exp\left(-\frac{\beta}{\theta}\right)$
- **Determine the MAP Estimator for θ**

Last Time

■ MAP Estimator Example

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- Assume inverse gamma prior $p(\theta) = \frac{\beta^\alpha}{\Gamma(\alpha)} \theta^{-\alpha-1} \exp\left(-\frac{\beta}{\theta}\right)$
- **Determine the MAP Estimator for θ**

$$\begin{aligned} p(x|\theta)p(\theta) &= \frac{1}{\sqrt{2\pi\theta}} \exp\left(\frac{-1}{2\theta} x^2\right) \frac{\beta^\alpha}{\Gamma(\alpha)} \theta^{-\alpha-1} \exp\left(-\frac{\beta}{\theta}\right) \\ &= \frac{1}{\sqrt{2\pi\theta}} \frac{\beta^\alpha}{\Gamma(\alpha)} \theta^{-\alpha-1} \exp\left(-\frac{1}{\theta} \left(\beta + \frac{x^2}{2}\right)\right) \end{aligned}$$

$$\ln(p(x|\theta)p(\theta)) = \frac{-1}{2} \ln(2\pi\theta) - (\alpha + 1) \ln(\theta) - \frac{1}{\theta} \left(\beta + \frac{x^2}{2}\right) + \ln\left(\frac{\beta^\alpha}{\Gamma(\alpha)}\right)$$

$$\frac{\partial \ln(p(x|\theta)p(\theta))}{\partial \theta} = \frac{-1}{2\theta} - (\alpha + 1) \frac{1}{\theta} + \frac{1}{\theta^2} \left(\beta + \frac{x^2}{2}\right) = 0$$

Last Time

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- **Determine the MAP Estimator for θ**

$$\frac{1}{\theta^2} \left(\beta + \frac{x^2}{2} \right) = \frac{1}{\theta} \left((\alpha + 1) + \frac{1}{2} \right)$$

$$\frac{1}{\theta^2} \left(\beta + \frac{x^2}{2} \right) = \frac{1}{\theta} \left(\alpha + \frac{3}{2} \right)$$

$$\hat{\theta} = \frac{\left(\beta + \frac{x^2}{2} \right)}{\left(\alpha + \frac{3}{2} \right)}$$

Bayesian Estimation: Linear MMSE Estimators

Linear MMSE Estimator

■ A simplification

- Assume the estimator is a linear system
- For a **single parameter** estimator

$$\hat{\theta} = \sum_{n=1}^N h_n x_n + a = \mathbf{h}^T \mathbf{x} + a$$

- For a **multi-parameter** estimator

$$\hat{\boldsymbol{\theta}} = \mathbf{H}\mathbf{x} + \mathbf{a}$$

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Linear MMSE Estimator

- **Goal:** To minimize the Bayesian Mean Square Error
 - **Step one:** Find the optimal estimate for a

Linear MMSE Estimator

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- **Step one:** Find the optimal estimate for a

$$BMSE = E \left[(\theta - \hat{\theta})^2 \right]$$

$$BMSE = E \left[(\theta - (\mathbf{h}^T \mathbf{x} + a))^2 \right]$$

$$\frac{\partial(BMSE)}{\partial a} = -2E[\theta - (\mathbf{h}^T \mathbf{x} + a)] = 0$$

$$a^* = E[\theta] - \mathbf{h}^T E[\mathbf{x}]$$

Linear MMSE Estimator

- **Goal:** To minimize the Bayesian Mean Square Error
 - **Step two:** Find the optimal estimate for $\mathbf{h}$

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$$BMSE = E \left[(\theta - (\mathbf{h}^T \mathbf{x} + E[\theta] - \mathbf{h}^T E[\mathbf{x}]))^2 \right]$$

$$BMSE = E \left[((\theta - E[\theta]) - (\mathbf{h}^T \mathbf{x} - \mathbf{h}^T E[\mathbf{x}]))^2 \right]$$

$$BMSE = E \left[((\theta - E[\theta]) - \mathbf{h}^T (\mathbf{x} - E[\mathbf{x}]))^2 \right]$$

$$BMSE = E \left[(\theta - E[\theta])^2 - \mathbf{h}^T (\mathbf{x} - E[\mathbf{x}]) (\theta - E[\theta]) \right]$$

Linear MMSE Estimator

- **Goal:** To minimize the Bayesian Mean Square Error

- **Step two:** Find the optimal estimate for $\mathbf{h}$

$$BMSE = C_{\theta\theta} - \mathbf{h}^T \mathbf{C}_{x\theta} - \mathbf{C}_{\theta x} \mathbf{h} + \mathbf{h}^T \mathbf{C}_{xx} \mathbf{h}$$

$$\frac{\partial(BMSE)}{\partial \mathbf{h}} = -2\mathbf{C}_{x\theta} + 2\mathbf{C}_{xx} \mathbf{h} = 0$$

$$\begin{aligned} \mathbf{C}_{xx} \mathbf{h} &= \mathbf{C}_{x\theta} \\ \mathbf{h}^* &= \mathbf{C}_{xx}^{-1} \mathbf{C}_{x\theta} \end{aligned}$$

Linear MMSE Estimator

- **Goal:** To minimize the Bayesian Mean Square Error

- **Step three:** Find the optimal estimator

$$\hat{\theta} = \mathbf{h}^T \mathbf{x} + a$$

$$a^* = E[\theta] - \mathbf{h}^T E[\mathbf{x}]$$

$$\mathbf{h}^* = \mathbf{C}_{xx}^{-1} \mathbf{C}_{x\theta}$$

$$\hat{\theta} = \mathbf{C}_{xx}^{-1} \mathbf{C}_{x\theta}^T \mathbf{x} + E[\theta] - \mathbf{C}_{xx}^{-1} \mathbf{C}_{x\theta}^T E[\mathbf{x}]$$

$$\hat{\theta} = E[\theta] + \mathbf{C}_{xx}^{-1} \mathbf{C}_{x\theta}^T (\mathbf{x} - E[\mathbf{x}])$$

$$BMSE = C_{\theta\theta} - \mathbf{h}^T \mathbf{C}_{x\theta} - \mathbf{C}_{\theta x} \mathbf{h} + \mathbf{h}^T \mathbf{C}_{xx} \mathbf{h}$$

$$BMSE = C_{\theta\theta} - 2\mathbf{C}_{x\theta} \mathbf{C}_{xx}^{-1} \mathbf{C}_{x\theta} + \mathbf{C}_{x\theta} \mathbf{C}_{xx}^{-1} \mathbf{C}_{xx} \mathbf{C}_{xx}^{-1} \mathbf{C}_{x\theta}$$

$$BMSE = C_{\hat{\theta}} = C_{\theta\theta} - \mathbf{C}_{x\theta} \mathbf{C}_{xx}^{-1} \mathbf{C}_{x\theta}$$

Have we seen these results before?

Linear MMSE Estimator

- **Goal:** To minimize the Bayesian Mean Square Error
 - **Step three:** Find the optimal estimator

$$\hat{\theta} = E[\theta] + \mathbf{C}_{xx}^{-1} \mathbf{C}_{x\theta}^T (\mathbf{x} - E[\mathbf{x}])$$
$$BMSE = C_{\hat{\theta}} = C_{\theta\theta} - \mathbf{C}_{x\theta} \mathbf{C}_{xx}^{-1} \mathbf{C}_{x\theta}$$

Have we seen these results before?

Linear MMSE Estimator

■ A simplification

- Assume the estimator is a linear system
- For a **single parameter** estimator

$$\hat{\theta} = \sum_{n=1}^N h_n x_n + a = \mathbf{h}^T \mathbf{x} + a$$

- For a **multi-parameter** estimator

$$\hat{\boldsymbol{\theta}} = \mathbf{H}\mathbf{x} + \mathbf{a}$$

Linear MMSE Estimator

- **Goal:** To minimize the Bayesian Mean Square Error

- The LMMSE Estimator is

$$\begin{aligned}\hat{\theta} &= Hx + a \\ a^* &= E[\theta] - HE[x] \\ H^* &= C_{xx}^{-1}C_{x\theta}\end{aligned}$$

$$\begin{aligned}\hat{\theta} &= E[\theta] + C_{xx}^{-1}C_{x\theta}^T(x - E[x]) \\ C_{\hat{\theta}} &= C_{\theta\theta} - C_{x\theta}C_{xx}^{-1}C_{x\theta}\end{aligned}$$

Have we seen these results before?

Linear MMSE Estimator

- **Goal: To minimize the Bayesian Mean Square Error**

- When we have a linear statistical model: $\mathbf{x} = \mathbf{G}\boldsymbol{\theta} + \mathbf{w}$, for noise $\mathbf{w}$
- We assume the expected value and covariance of $\boldsymbol{\theta}$ and $\mathbf{w}$ are both known.

$$\begin{aligned}\hat{\boldsymbol{\theta}} &= E[\boldsymbol{\theta}] + (\mathbf{C}_{\boldsymbol{\theta}\boldsymbol{\theta}} + \mathbf{G}^T \mathbf{C}_{\mathbf{w}}^{-1} \mathbf{G})^{-1} \mathbf{G}^T \mathbf{C}_{\mathbf{w}}^{-1} (\mathbf{x} - \mathbf{G}E[\mathbf{x}]) \\ \mathbf{C}_{\hat{\boldsymbol{\theta}}} &= \mathbf{C}_{\boldsymbol{\theta}\boldsymbol{\theta}} - \mathbf{G}^T \mathbf{C}_{\mathbf{w}}^{-1} \mathbf{G}\end{aligned}$$

Have we seen these results before?

Bayesian Estimation: Linear MMSE Example

Last Time

■ LMMSE Estimator Example

- Gaussian likelihood $x|\theta \sim \mathcal{N}(0, \sigma^2)$
- **Determine the LMMSE Estimator for $\theta = x^2$ (Chi-squared prior)**

Last Time

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- Gaussian likelihood $x|\theta \sim \mathcal{N}(0, \sigma^2)$
- **Determine the LMMSE Estimator for $\theta = x^2$ (Chi-squared prior)**

$$E[(\theta - \hat{\theta})^2] = E[(\theta - a_0x - a_1)^2]$$

$$\frac{\partial E[(\theta - \hat{\theta})^2]}{\partial a_1} = E[2(\theta - a_0x - a_1)] = 0$$

$$E[\theta - a_0x - a_1] = 0$$

$$E[\theta] - a_1 = 0$$

$$a_1 = E[\theta] = \sigma^2$$

Last Time

■ LMMSE Estimator Example

- Gaussian likelihood $x|\theta \sim \mathcal{N}(0, \sigma^2)$
- **Determine the LMMSE Estimator for $\theta = x^2$ (Chi-squared prior)**

$$E[(\theta - \hat{\theta})^2] = E[(\theta - a_0x - a_1)^2]$$

$$\frac{\partial E[(\theta - \hat{\theta})^2]}{\partial a_0} = E[(\theta - a_0x - a_1)x] = 0$$

$$E[\theta x - a_0x^2 - a_1x] = 0$$

$$E[\theta x] - a_0E[x^2] = 0$$

$$E[x^3] - a_0E[x^2] = 0$$

$$a_0 = 0$$

Last Time

■ LMMSE Estimator Example

- Gaussian likelihood $x|\theta \sim \mathcal{N}(0, \sigma^2)$
- **Determine the LMMSE Estimator for $\theta = x^2$ (Chi-squared prior)**

$$\hat{\theta} = \sigma^2$$

$$\begin{aligned} BMSE &= E \left[(\theta - \hat{\theta})^2 \right] = E[(x^2 - \sigma^2)^2] \\ &= E[x^4 - 2x^2\sigma^2 + \sigma^4] \\ &= 3\sigma^4 - 2\sigma^4 + \sigma^4 \\ &= 2\sigma^4 \end{aligned}$$

Last Time

■ LMMSE Estimator Example

- Gaussian likelihood $x|\theta \sim \mathcal{N}(0, \sigma^2)$

- In contrast, the estimator:

$$\hat{\theta} = x^2$$

$$\begin{aligned} BMSE &= E \left[(\theta - \hat{\theta})^2 \right] = E[(x^2 - x^2)^2] \\ &= 0 \end{aligned}$$

Bayesian Estimation: Multiple data example

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■ MMSE / MAP / LMMSE Estimator Example

- Consider jointly Gaussian random variables θ, x_1, x_2 . $\mathbf{x} = \begin{bmatrix} x_1 \\ x_2 \end{bmatrix}$
- Assume x_1, x_2 are independent.
- **Determine the MMSE / MAP / LMMSE Estimator for θ .**

Last Time

■ MMSE / MAP / LMMSE Estimator Example

- Consider jointly Gaussian random variables θ, x_1, x_2 . $\mathbf{x} = \begin{bmatrix} x_1 \\ x_2 \end{bmatrix}$
- Assume x_1, x_2 are independent.
- Determine the MMSE / MAP / LMMSE Estimator for θ .

■ MMSE Estimator: $\hat{\theta} = E[\theta | x_1, x_2] = E[\theta] + \mathbf{C}_{\theta x} \mathbf{C}_{xx}^{-1} (\mathbf{x} - E[\mathbf{x}])$

$$\begin{aligned} \mathbf{C}_{\theta x} &= E[\theta \mathbf{x}^T] = [\mathbf{C}_{\theta x_1} \quad \mathbf{C}_{\theta x_2}] \\ \mathbf{C}_{xx} &= E[\mathbf{x} \mathbf{x}^T] = \begin{bmatrix} E[x_1 x_1^T] & E[x_1 x_2^T] \\ E[x_1 x_2^T] & E[x_2 x_2^T] \end{bmatrix} = \begin{bmatrix} C_{x_1 x_1} & 0 \\ 0 & C_{x_2 x_2} \end{bmatrix} \end{aligned}$$

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■ MMSE / MAP / LMMSE Estimator Example

- Consider jointly Gaussian random variables θ, x_1, x_2 . $\mathbf{x} = \begin{bmatrix} x_1 \\ x_2 \end{bmatrix}$
- Assume x_1, x_2 are independent.
- Determine the MMSE / MAP / LMMSE Estimator for θ .

■ MMSE Estimator: $\hat{\theta} = E[\theta | x_1, x_2] = E[\theta] + \mathbf{C}_{\theta x} \mathbf{C}_{xx}^{-1} (\mathbf{x} - E[\mathbf{x}])$

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$$\hat{\theta} = E[\theta] + \mathbf{C}_{\theta x_1} \mathbf{C}_{x_1 x_1}^{-1} (x_1 - E[x_1]) + \mathbf{C}_{\theta x_2} \mathbf{C}_{x_2 x_2}^{-1} (x_2 - E[x_2])$$

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■ MMSE / MAP / LMMSE Estimator Example

- Consider jointly Gaussian random variables θ, x_1, x_2 . $\mathbf{x} = \begin{bmatrix} x_1 \\ x_2 \end{bmatrix}$
- Assume x_1, x_2 are independent.
- Assume θ is zero mean.
- Determine the MMSE / MAP / LMMSE Estimator for θ .

■ MMSE Estimator: $\hat{\theta} = E[\theta|x_1, x_2] = E[\theta] + \mathbf{C}_{\theta\mathbf{x}}\mathbf{C}_{\mathbf{x}\mathbf{x}}^{-1}(\mathbf{x} - E[\mathbf{x}])$

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Additive property of independent data sets

Bayesian Estimation:

Dynamic / state models (scalar case)

Kalman Filters

■ Consider the signal problem

$$\begin{aligned}\mathbf{x} &= \mathbf{s} + \mathbf{w} \\ x[n] &= s[n] + w[n]\end{aligned}$$

where $\mathbf{w}$ is Gaussian noise with covariance $\mathbf{I}\sigma^2$.

What is the MVU Estimator of s ?

Why is this a pretty terrible estimator?

How could we improve it?

Kalman Filters

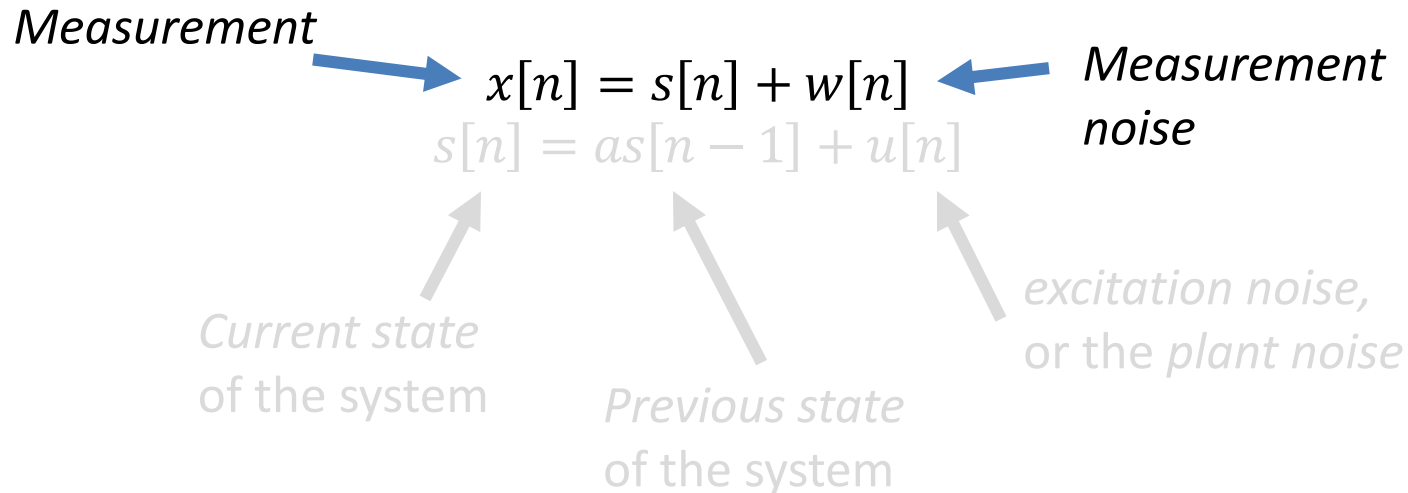
■ The Idea:

- The signal has:
 - Nonstationary statistics (i.e., the mean changes over time)
 - Correlated values ($x[n]$ is related to $x[n-1]$)
- Design an estimator such that
 - The current estimated value based on previous measurements
 - This is the idea of a ***filter***.

Kalman Filters

■ First-order Gauss-Markov *Process*:

- Assume our signal can be represented by a *dynamical* or *state model*

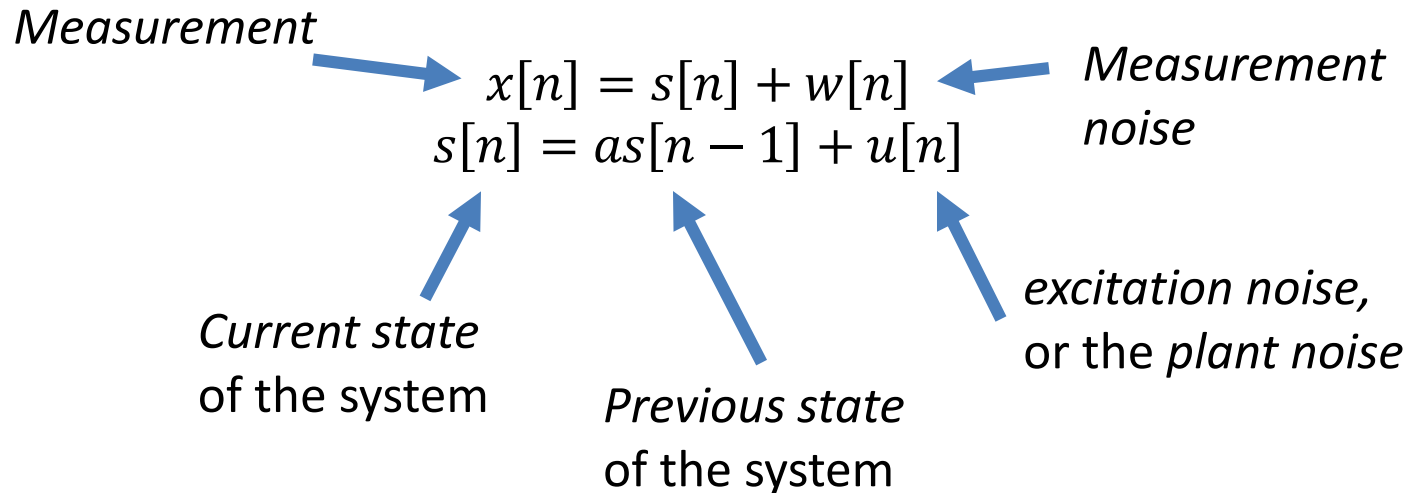


This is our current model.

Kalman Filters

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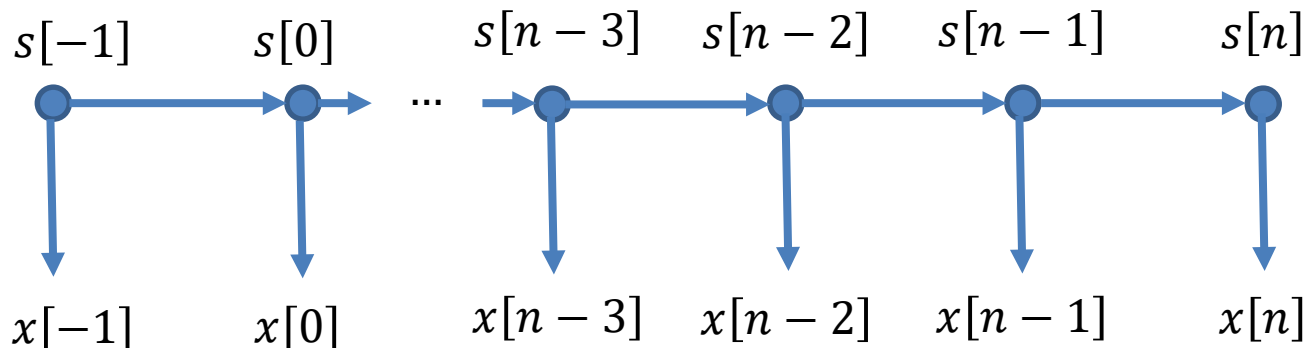
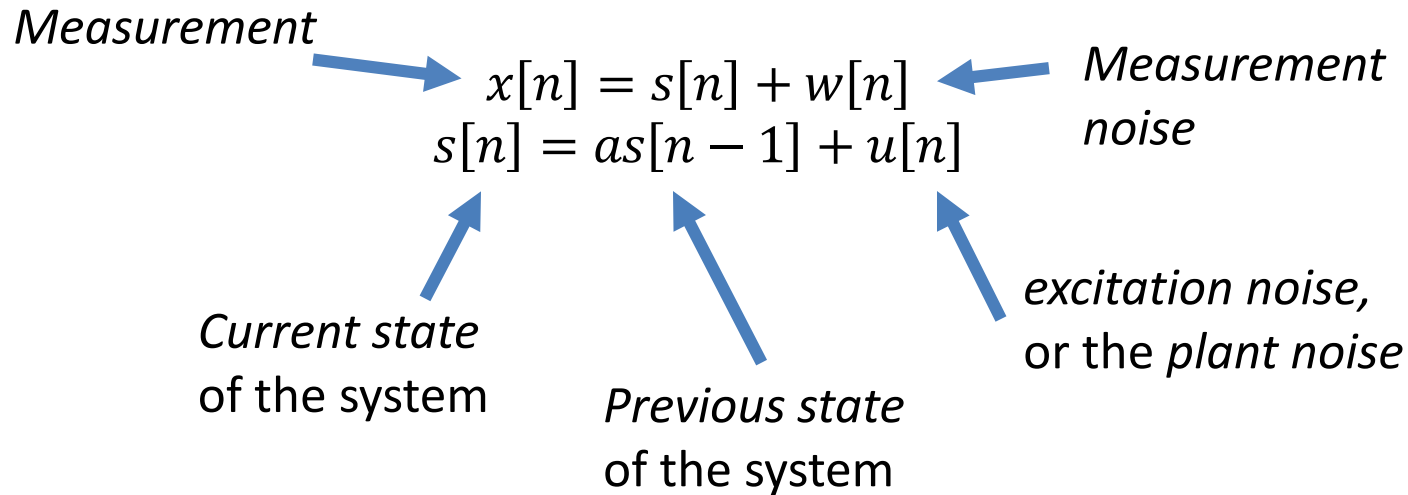


This is our has an additional state model.

Kalman Filters

■ First-order Gauss-Markov *Process*:

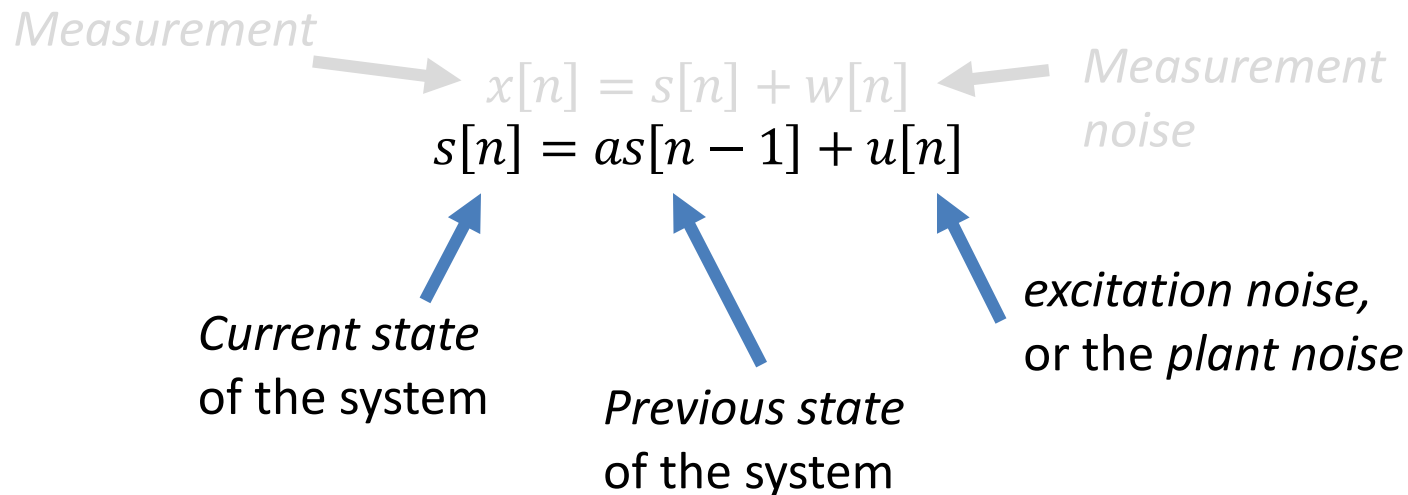
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Kalman Filters

■ First-order Gauss-Markov *Process*:

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What the closed form solution for $s[n]$?

Kalman Filters

■ First-order Gauss-Markov *Process*:

- Assume our signal can be represented by a *dynamical* or *state model*

$$s[n] = as[n - 1] + u[n]$$

$$s[0] = as[-1] + u[0]$$

$$s[1] = as[0] + u[1] = a^2s[-1] + au[0] + u[1]$$

$$s[2] = a^3s[-1] + a^2u[0] + au[1] + u[2]$$

$$s[3] = a^4s[-1] + a^3u[0] + a^2u[1] + au[2] + u[3]$$

Kalman Filters

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$$s[3] = a^4s[-1] + a^3u[0] + a^2u[1] + au[2] + u[3]$$

$$s[n] = a^{n+1}s[-1] + \sum_{k=0}^n a^k u[n-k]$$

Kalman Filters

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- Assume our signal can be represented by a *dynamical* or *state model*

$$s[n] = a^{n+1}s[-1] + \sum_{k=0}^n a^k u[n-k]$$

- Assume (for now)
 - $s[-1] \sim \mathcal{N}(\mu_s, \sigma_s^2)$
 - $u[n] \rightarrow \mathbf{u} \sim \mathcal{N}(\mathbf{0}, \sigma_u^2 \mathbf{I})$
 - $s[-1]$ and $u[n]$ are independent

What is the expected value of $s[n]$?

Kalman Filters

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What is the expected value of $s[n]$?

$$E[s[n]] = a^{n+1}\mu_s$$

Kalman Filters

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What is the expected value of $s[n]$ (in terms of $s[n-1]$)?

Kalman Filters

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What is the expected value of $s[n]$ (in terms of $s[n-1]$)?

$$E[s[n]] = aE[s[n-1]]$$

Kalman Filters

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 - $s[-1]$ and $u[n]$ are independent

What is the covariance between of $s[n], s[m]$?

Kalman Filters

■ First-order Gauss-Markov *Process*:

- Assume our signal can be represented by a *dynamical* or *state model*

What is the covariance between of $s[n]$, $s[m]$?

$$\begin{aligned} & E[(s[n] - a^{n+1}\mu_s)(s[m] - a^{m+1}\mu_s)] \\ &= E\left[\left(a^{n+1}s[-1] + \sum_{k=0}^n a^k u[n-k] - a^{n+1}\mu_s\right)\left(a^{m+1}s[-1] + \sum_{\ell=0}^m a^\ell u[m-\ell] - a^{m+1}\mu_s\right)\right] \\ &= E\left[\left(a^{n+1}(s[-1] - \mu_s) + \sum_{k=0}^n a^k u[n-k]\right)\left(a^{m+1}(s[-1] - \mu_s) + \sum_{\ell=0}^m a^\ell u[m-\ell]\right)\right] \end{aligned}$$

Due to the independence of $s[-1]$ and $u[n]$

$$= a^{m+n+2}\sigma_s^2 + \sum_{k=0}^n \sum_{\ell=0}^m a^{k+\ell} E[u[m-k]u[n-\ell]]$$

Due to the independence of each $u[n]$

$$C(s[n], s[m]) = a^{m+n+2}\sigma_s^2 + \sigma_u^2 a^{m-n} \sum_{k=0}^n a^{2k}$$

Kalman Filters

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- Assume our signal can be represented by a *dynamical* or *state model*

What is the covariance between of $s[n], s[m]$?

$$C(s[n], s[m]) = a^{m+n+2} \sigma_s^2 + \sigma_u^2 a^{m-n} \sum_{k=0}^n a^{2k}$$

$$\text{var}(s[n]) = C(s[n], s[n]) = a^{2n+2} \sigma_s^2 + \sigma_u^2 \sum_{k=0}^n a^{2k}$$

If $a = 1$?:

Kalman Filters

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$$C(s[n], s[n]) = \sigma_s^2 + n \sigma_u^2$$

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If $|a| < 1, n \rightarrow \infty$: ?

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If $|a| < 1, n \rightarrow \infty$: (use geometric series relationship)

$$C(s[n], s[m]) = \frac{\sigma_u^2 a^{m-n}}{1 - a^2}$$

$$C(s[n], s[n]) = \frac{\sigma_u^2}{1 - a^2}$$

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In terms of $s[n-1]$:

$$\text{var}(s[n-1]) = a^{2n}\sigma_s^2 + \sigma_u^2 \sum_{k=0}^{n-1} a^{2k}$$

$$\text{var}(s[n]) = a^2 \text{var}(s[n-1]) + \sigma_u^2$$

Kalman Filters

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$$s[n] = as[n - 1] + u[n]$$

$$E[s[n]] = aE[s[n - 1]]$$
$$\text{var}(s[n]) = a^2\text{var}(s[n - 1]) + \sigma_u^2$$

